Journal of Modern Applied Statistical Methods

Volume 7 | Issue 2 Article 11

11-1-2008

Frequency Domain Modeling with Piecewise Constant Spectra

Erhard Reschenhofer *University of Vienna*, erhard.reschenhofer@univie.ac.at

Part of the <u>Applied Statistics Commons</u>, <u>Social and Behavioral Sciences Commons</u>, and the <u>Statistical Theory Commons</u>

Recommended Citation

Reschenhofer, Erhard (2008) "Frequency Domain Modeling with Piecewise Constant Spectra," *Journal of Modern Applied Statistical Methods*: Vol. 7: Iss. 2, Article 11.

DOI: 10.22237/jmasm/1225512600

Frequency Domain Modeling with Piecewise Constant Spectra

Erhard Reschenhofer University of Vienna, Austria

Using piecewise constant functions as models for the spectral density of the differenced log real U.S. GDP it was found that these models have the capacity to compete with the spectral densities implied by ARMA models. According to AIC and BIC the piecewise constant spectral densities are superior to ARMA.

Key words: Spectral analysis, piecewise constant spectra, ARMA spectra, aggregate output.

Introduction

Univariate ARMA models are used in empirical economics as simple, purely statistical models for properly transformed macroeconomic time series (such as the first differences of the logs of the real GDP), and for the description of the serial correlation in the errors of more complex models such as linear or nonlinear multivariate regression models. A typical example of the first type is the study by Campbell & Mankiw (1987) who used ARMA(p,q) models with $p \le 3$ and q \leq 3 to investigate the long-run behavior of aggregate output. The persistence of output shocks can be measured by the cumulative impulse response or, equivalently, by the value of the spectral density at frequency zero, however, two drawbacks exist. The first is that the model parameters must be estimated by numerical optimization routines, which depend heavily on the starting values and can easily get stuck at local optima (e.g., Hauser, et al., 1999). The second is the extreme sensitivity of inference to the order of the ARMA representation (e.g., Christiano & Eichenbaum, 1990).

Recently, interest has shifted from univariate to multivariate modeling (e.g., Blanchard & Quah,

Erhard Reschenhofer is Associate Professor in the Department of Statistics and Decision Support Systems. Email: erhard.reschenhofer@univie.ac.at.

1989; Pesaran, et al., 1993; Pesaran & Shin, . However, a multivariate approach based on economic theory and the information contained in a much larger data set is not necessarily better than a simple univariate time series model, because both the estimation identification of multivariate models is many orders of magnitude more difficult. But even in situations where multivariate models outperform univariate models, the latter are often used as benchmarks for the former (see, Schumacher & Dreger, 2004). Thus, univariate ARMA models still have an important role to This article proposes competitive alternatives to ARMA models for the purpose of densities estimating the spectral macroeconomic time series.

Methodology

The following piecewise constant functions are proposed:

$$\begin{split} g_r(\omega) &= a(b_1 \mathbf{1}_{[\alpha_0,\alpha_1)} + b_2 \mathbf{1}_{[\alpha_1,\alpha_2)} + \dots \\ &+ b_{r\text{-}1} \mathbf{1}_{[\alpha_{r-2},\alpha_{r-1})} + \mathbf{1}_{[\alpha_{r-1},\alpha_r]}), \ \omega \in [0,\pi], \end{split}$$

where $r \ge 2$ and $0=\alpha_0<\alpha_1<...<\alpha_r=\pi$, for the approximation of the spectral densities of macroeconomic time series. There are 2(r-1)+1 parameters that must be estimated, $a, b_1, ..., b_{r-1}, \alpha_1, ..., \alpha_{r-1}$. An obvious choice for the first parameter is:

FREQUENCY DOMAIN MODELING WITH PIECEWISE CONSTANT SPECTRA

$$a = \frac{I_1 + ... + I_m}{b_1 s_1 + b_2 (s_2 - s_1) + ... + b_{r-1} (s_{r-1} - s_{r-2}) + (m - s_{r-1})},$$
(2)

where s_i is the largest integer such that:

$$\frac{2\pi s_j}{n} < \alpha_j. \tag{3}$$

The parameters $b_1, ..., b_{r-1}, ..., s_1, ..., s_{r-1}$ can be found by maximizing the Whittle likelihood

$$\prod_{k=1}^{m} \frac{1}{g_r(\omega_k)} \exp\left(-\frac{I_k}{g_r(\omega_k)}\right), \quad (4)$$

or, equivalently,

$$-\sum_{k=1}^{n} \log(g_{r}(\omega_{k})) - \sum_{k=1}^{m} \frac{I_{k}}{g_{r}(\omega_{k})}, \quad (5)$$

where

$$\omega_{k} = \frac{2\pi k}{n}, k=1,...,m.$$
 (6)

The parameters $\alpha_1, ..., \alpha_{r-1}$ can be obtained from $s_1, ..., s_{r-1}$ via

$$\alpha_{j} = \frac{2\pi s_{j}}{n} + \frac{\pi}{n} \,. \tag{7}$$

To demonstrate the usefulness of this approach, the seasonally adjusted quarterly real U.S. GDP from 1947.1 to 2007.1 was downloaded from FRED® (Federal Reserve Economic Data) and the spectral density of the first differences of the log GDP was approximated by the piecewise constant functions g_j , j=2,3,4.

Results

Figure 1 compares the three piecewise constant spectral densities with the best three ARMA spectral densities selected by BIC. One of these three ARMA models, namely the ARMA(3,2) model, is the best ARMA model according to AIC. Apart from the ARMA models of order (2,3) and (3,3), whose spectral densities are very similar to that of the ARMA(3,2) model, all other ARMA models $(p \le 8 \& q = 0, p = 0 \& q \le 8, 1 \le p, q \le 3)$ have much higher AIC values than the ARMA(3,2) model. To facilitate the comparison between the piecewise constant spectral densities g2, g3, and g₄, and the ARMA spectral densities slightly modified AIC and BIC values (AIC* and BIC*) obtained from the Whittle likelihood were used. Among the top models both according to AIC* and BIC* (see Tables 1 and 2) are g₂, g₃, and g₄. Overall, g₂ has the smallest BIC* value and g₄ has the smallest AIC* value.

Table 1: AIC values (obtained from the Whittle likelihood) for piecewise constant spectral densities g(r) & ARMA(p,q) spectral densities, respectively, fitted to the differenced log real U.S. GDP

	1	2	3	4	5	6	7	8
g(r)		-2456.8	-2457.5	-2461.6				
AR(p)	-2447.3	-2446.9	-2447.7	-2448.6	-2448.1	-2446.2	-2444.2	-2442.4
MA(q)	-2440.5	-2448.2	-2447.8	-2445.9	-2447.7	-2445.7	-2443.8	-2445.2
ARMA(1,q)	-2446.2	-2447.2	-2445.8					
ARMA(2,q)	-2445.8	-2450.7	-2457.6					
ARMA(3,q)	-2449.1	-2459.2	-2457.1					

RESCHENHOFER

Table 2: BIC values (obtained from the Whittle l	likelihood) for piecewise constant spectral densities g(r) &
ARMA(n a) spectral densities respecti	tively fitted to the differenced log real U.S. GDP

	1	2	3	4	5	6	7	8
g(r)		-2446.3	-2440.1	-2437.3				
AR(p)	-2440.3	-2436.5	-2433.7	-2431.2	-2427.3	-2421.9	-2416.4	-2411.1
MA(q)	-2433.5	-2437.8	-2433.9	-2428.5	-2426.8	-2421.3	-2416.0	-2413.9
ARMA(1,q)	-2435.8	-2433.3	-2428.4					
ARMA(2,q)	-2431.9	-2433.3	-2436.8					
ARMA(3,q)	-2431.7	-2438.3	-2432.8					

Conclusion

The results obtained show that piecewise constant spectral densities are extremely useful tools for the spectral analysis of macroeconomic time series and can outperform the more sophisticated ARMA spectral densities. This finding is striking given that twenty-five ARMA spectral densities were tried but only three piecewise constant spectral densities. It may also serve as a severe warning not to over-interpret certain characteristics of estimated ARMA spectral densities such as a decline or incline near frequency zero.

References

Blanchard, J. O., & Quah, D. (1989). The dynamic effects of aggregate supply and demand disturbances. *American Economic Review*, 79, 655-673.

Campbell, J. Y., & Mankiw, N. G. (1987). Are output fluctuations transitory? *The Quarterly Journal of Economics*, *102*, 857-880.

Christiano, L. J., & Eichenbaum, M. (1990). Unit roots in real GNP: Do we know, and do we care? *Carnegie-Rochester Conference Series on Public Policy*, 32, 7-61.

Hauser, M. A., Pötscher, B. M., & Reschenhofer, E. (1999). Measuring persistence in aggregate output: ARMA models, fractionally integrated ARMA models and nonparametric procedures. *Empirical Economics*, *24*, 243-269.

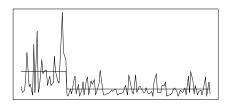
Pesaran, H. M., Pierse R. G., & Lee, K. C. (1993). Persistence, co-integration and aggregation: A disaggregate analysis of output fluctuations in the U.S. economy. *Journal of Econometrics*, 56, 57-88.

Pesaran, H. M., & Shin, Y. (1996). Cointegration and speed of convergence to Equilibrium. *Journal of Econometrics*, 71, 117-143.

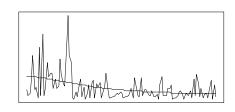
Schumacher, C., & Dreger, C. (2004). Estimating large-scale factor models for economic activity in Germany: Do they outperform simpler models? *Journal of Economics and Statistics* 224, 732-750.

Figure 1: Periodogram of differenced log GDP together with piecewise constant spectral densities (with two, three, and four pieces) & ARMA spectral densities AR(1), MA(2), ARMA(3,2)

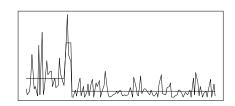




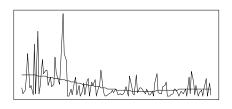
AIC*=-2447.3, BIC*=-2440.3



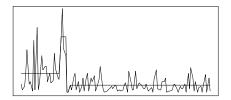
AIC*=-2457.5, BIC*=-2440.1



AIC*=-2448.2, BIC*=-2437.8



AIC*=-2461.6, BIC*=-2437.3



AIC*=-2459.2, BIC*=-2438.3

