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# Discriminating Between Generalized Exponential Distribution and Some Life Test Models Based on Population Quantiles

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# **Discriminating Between Generalized Exponential Distribution and Some Life Test Models Based on Population Quantiles**

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A test statistic based on population quantiles using sample order statistics is suggested. The quantiles of the test statistics are evaluated for generalized exponential distribution. Similar test statistic based on moments of sample order statistic is referred and the proposed test formula is compared with it. Between the pairs of the above models it is established that the test formula proposed by us is more effective and useful than the formula based on the moments of order statistics as developed by Sultan [\(2007\)](#page-8-0).

*Keywords:* Population quantiles, generalized exponential distribution

### **Introduction**

The three-parameter generalized exponential (GE) distribution has its probability density function (pdf) as

$$
f(x) = \frac{\alpha}{\sigma} \left( 1 - e^{-\left(\frac{x-\mu}{\sigma}\right)} \right)^{\alpha-1} e^{-\left(\frac{x-\mu}{\sigma}\right)}, x > \mu, \mu > 0, \alpha > 0, \sigma > 0 \tag{1}
$$

Its cumulative distribution function (cdf) is given by

<span id="page-1-0"></span>
$$
F(x) = \frac{\alpha}{\sigma} \left( 1 - e^{-\left(\frac{x-\mu}{\sigma}\right)} \right)^{\alpha}, x > \mu, \mu > 0, \alpha > 0, \sigma > 0
$$
 (2)

The two-parameter GE distribution has its pdf as

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$$
f(x) = \frac{\alpha}{\sigma} \left( 1 - e^{-\left(\frac{x}{\sigma}\right)} \right)^{\alpha - 1} e^{-\left(\frac{x}{\sigma}\right)}, x > 0, \alpha > 0, \sigma > 0 \tag{3}
$$

Its cdf is given by

<span id="page-2-0"></span>
$$
F(x) = \frac{\alpha}{\sigma} \left( 1 - e^{-\left(\frac{x}{\sigma}\right)} \right)^{\alpha}, x > 0, \alpha > 0, \sigma > 0
$$
 (4)

The GE distribution was introduced by Gupta and Kundu [\(1999\)](#page-7-0). It is compared with the two-parameter Gamma and Weibull distributions in Gutpa and Kundu [\(2001a\)](#page-7-1). Different models of estimations are discussed in Gutpa and Kundu [\(2001b\)](#page-7-2). Raqab and Ahsanullah [\(2001\)](#page-8-1) and Raqab [\(2002\)](#page-8-2) studied the properties of order and record statistics from the two-parameter GE distribution respectively. Discriminating between gamma and GE distribution were studied by Gutpa and Kundu [\(2004\)](#page-7-3). Discriminating between lognormal and GE distribution was given in Kundu et al [\(2005\)](#page-7-4). The expected values of order statistics may not always be available in numerical form nor analytically simple beyond a given sample size. However if the distribution function is invertible analytically the population quantile for any 'n' can be easily obtained. Also moment of order statistics are conceptually similar to the population quantiles with an admissible measure of closeness. Therefore, quantiles are used to develop the test statistic and to distinguish the GE distribution from other well-known life testing models. The proposed work is similar to that of Sultan [\(2007\)](#page-8-0) wherein moments of order statistics are used to develop the test statistic with GE distributions null population. The aim of this article is to explore the usefulness of analytical expressions of population quantiles of GE distribution. In section 2 the GE distribution and its quantiles are developed. In section 3 the goodness of fit tests of the two-parameter and three parameter GE distribution are developed. Section 4 deals with the power of the proposed test procedure in comparison with that of Sultan [\(2007\)](#page-8-0) with the same alternative populations. In section 5 the performance of quantiles of GE distribution is tested, and Section 6 contains concluding remarks.

#### **The GE distribution and its quantiles**

The p<sup>th</sup> quantile of population is defined as the solution of the equation  $F(x) = p_i$ where  $F(x)$  is the cdf given in (1.4). This is also called the standard population

quantile. If  $x_1, x_2, \ldots, x_n$  is an ordered sample of size n and  $p_i = \frac{i}{n+1}$  then the solution of  $F(x) = p_i$  is defined as i<sup>th</sup> population quantile corresponding to its order statistic  $x_i$ . In the sample this is denoted by  $d_i$  i.e.,  $F(d_i) = \frac{i}{n+1}$ . Expected value of its order statistic in the sample is denoted by  $\mu_i$ , the theory of order statistics indicate that  $\mu_i$ ,  $d_i$  can be approximated by each other. If the distribution function of the population is in a closed form,  $d_i$ 's sometimes can be obtained more easily than  $\alpha$ 's, moments of order statistics. This possibility is explored in developing the proposed test statistics of this article. Given the form of  $F(x)$  and a natural number n, d<sub>i</sub>'s can be obtained by inverting the population distribution function. This was done for the GE distribution with the shape parameter  $\alpha = 0.5, 2.0$  and sample size  $n = 10, 20, 25$ . To make use of them in the proposed test statistic, the details are given in the following sections.

#### **Goodness-of-Fit Test using quantiles**

**Test for two-parameter case** Let  $x_1, x_2, \ldots, x_n$  denote a sample from twoparametric GE distribution. The correlation type goodness of fit test procedure in this case using quantiles can be formed as follows:

 $H_0$ : *F* is correct, that is  $x_1, x_2, \dots, x_n$  have  $GE(0, \sigma, \alpha)$  given in [\(4\)](#page-2-0) versus  $H_1$ : *F* is not correct, that is  $x_1, x_2, \ldots, x_n$  have another cdf and the test statistic used to run the test is given by

$$
T_1 = \frac{\sum_{i=1}^n x_i d_i}{\sqrt{\sum_{i=1}^n x_i^2 \sum_{i=1}^n d_i^2}}
$$
 (5)

The test statistic  $T_1$  is simulated through Monte-Carlo method based on 10,000 simulations. Table 1 represents the percentages points of  $T<sub>1</sub>$  for sample sizes  $n = 10, 20, 25$  and  $\alpha = 0.5, 2.0$ . It can be seen from the [Table 1,](#page-4-0) the percentage points of  $T<sub>1</sub>$  follow the naturally expected order.

	$\alpha$ n 0.50%	$1\%$	2%	$2.50\%$	5%	10%	20%	30%	40%	50%
	0.5 10 0.9947	0.9931	0.991	0.9901		0.9865 0.9808	0.9708   0.9611		0.9499 0.9374	
	20 0.9921	0.9903	0.9879 0.9865 0.9818 0.9747				0.963		0.9506  0.9375  0.9221	
	25 0.991	0.989	0.986	0.9847		0.9796  0.9717  0.959		0.9462 0.9334		0.9189
	10 0.9954	0.9944	0.9933  0.9928  0.9908  0.9877						0.9829 0.9782 0.9734 0.9674	
	20 0.9944		0.9932 0.9918 0.9912 0.9891 0.9863 0.9818 0.9773 0.9723 0.9665							
	25 0.9936	0.9925	0.9911	0.9906	0.9886	0.9856	0.9812	0.977	0.9722	0.9669

<span id="page-4-0"></span>**Table 1.** Percentage points of  $T_1$ 

*Test for the three-parameter case* Let  $x_1, x_2, \ldots, x_n$  denote a sample from threeparametric GE distribution and let  $Z_i = X_i - X_1$  and  $v_i = d_i - d_i$ ,  $i = 1, 2, \dots n - 1$ . The correlation type goodness of fit test in this case using quantiles can be formed as follows:

 $H_0$ : *F* is correct, that is  $x_1, x_2, \ldots, x_n$  have  $GE(0, \sigma, \alpha)$  given in [\(2\)](#page-1-0) versus  $H_1$ : *F* is not correct, that is  $x_1, x_2, \ldots, x_n$  have another cdf and the test statistic used to run the test is given by

$$
T_2 = \frac{\sum_{i=1}^n z_i v_i}{\sqrt{\sum_{i=1}^n z_i^2 \sum_{i=1}^n v_i^2}}
$$
(6)

The statistic  $T_2$  is simulated through Monte-Carlo method based on 10,000 simulations. Table 2 represents the percentages points of  $T_2$  for sample sizes  $n = 10, 20, 25$  and  $\alpha = 0.5, 2.0$ . It can be seen from the [Table 2](#page-4-1) the percentage points of  $T_2$  follow the naturally expected order.

<span id="page-4-1"></span>**Table 2.** Percentage points of  $T$ <sub>2</sub>

$\alpha$	n 0.50%	$1\%$	2% 2.50%	5%	10%	<b>20%</b>	$30\%$	40%	50%
							0.5 10 0.9951 0.9934 0.9914 0.9905 0.9869 0.9813 0.9712 0.9615 0.9505 0.9382		
		20 0.9923 0.9904 0.998			0.9865  0.9818  0.9748  0.963  0.507			0.9376 0.9222	
							25 0.9912 0.9891 0.9861 0.9847 0.9797 0.9718 0.959 0.9462 0.9334		0.919
							10 0.9968 0.9958 0.9947 0.9943 0.9928 0.9902 0.9857 0.9812 0.9767		0.971
		20 0.9953 0.9942 0.9929 0.9924 0.9905 0.9879 0.9834 0.979						0.974	0.9682
		25 0.9945 0.9935 0.9922 0.9917 0.9898 0.987 0.9826					0.9785	0.9737	0.9683

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#### **Power of the test**

The power of the test is calculated by replacing  $GE(\mu, \sigma, \alpha)$  random variates generator in the simulation program with generators from the alternative distributions including: normal, lognormal, Cauchy, Weibull and gamma. Based on different sample sizes and 10,000 simulations, the power is calculated to be

$$
Power = \frac{\text{\# of rejections of } H_0}{10,000}
$$

Where  $H_0$  is rejected if  $T_1(T_2)$  greater than or equal to the corresponding percentage points given in [Table](#page-4-0) 1 [\(Table 2](#page-4-1) and  $T_1(T_2)$ ) is evaluated from the alternative distributions. [Table 3](#page-5-0) and [4](#page-6-0) represent the power of the test for the twoparameter and three-parameter cases, respectively. The different alternative distributions considered are: (i) normal distribution  $N(\mu, \sigma)$ , (ii) lognormal Ln(μ, σ), (iii) Weibull distribution with location parameter µ, scale parameter σ and shape parameter  $\alpha$ , W( $\mu$ ,  $\sigma$ ,  $\alpha$ ), (iv) gamma distribution with location parameter  $\mu$ , scale parameter  $\sigma$  and shape parameter k  $G(\mu, \sigma, k)$  and (v0 Cauchy distribution with location parameter  $\mu$  and scale parameter  $\sigma C(\mu, \sigma)$ . [Table 3](#page-5-0) and [4](#page-6-0) indicate that the correlation test has good power to reject sample from the chosen alternative distributions.

		N(0,1)			W(0,1,3)	G(0,1,7)	
$\alpha$	n	5%	10%	5%	10%	5%	10%
0.5	10	0.9227	0.745	1	1	0.9991	0.9946
	20	0.9999	0.9985	1	1	0.9845	0.9306
	25	0.9989	0.9986	1	1	0.9959	0.9857
2	10	0.9971	0.9999	0.9991	0.9947	1	0.9996
	20	0.9997	0.9996	1	0.998	0.9972	0.9944
	25	0.9998	0.9997	1	1	0.9995	0.9986

<span id="page-5-0"></span>**Table 3.** Power of the test of the two-parameter case  $(\sigma = 1)$ 

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		LN(1,5)			W(0,1,6)	C(0,1)	
α	n	5%	10%	5%	10%	5%	10%
0.5	10	0.975	0.937	1	1		
	20	0.975	0.943	1			1
	25	0.982	0.959	1			1
$\overline{2}$	10	1		1	1		
	20						
	25			1			

<span id="page-6-0"></span>**Table 4.** Power of the test of the two-parameter case  $(\mu = 0, \sigma = 1)$ 

Tables similar to that of [3](#page-5-0) and [4](#page-6-0) are available in Sultan [\(2007\)](#page-8-0), evaluated using the moments of order statistics. By comparison, notice that the coverage probability given in the tables are uniformly larger than what are given in Sultan [\(2007\)](#page-8-0). Therefore, the test statistic proposed based on the population quantiles is more powerful than that based on the moments of order statistics. Moreover, for a distributional GE distribution moments of order statistics are not available completely beyond a given sample size whereas population quantiles are available for any sample size provided the mathematical form of the cdf is analytically invertible. Therefore it can be concluded that the proposed test statistic T is more powerful than that of Sultan [\(2007\)](#page-8-0).

#### **Numerical Examples**

In order to show the performances of the test of GE distribution in both cases (two-parameter and three-parameter), four sets of order statistics each of size 25 were simulated, they are

- 1. Sample from GE(0,1,2): two-parameter case of the GE distribution with scale parameter is equal to 1 and shape parameter is equal to 2
- 2. Sample from  $GE(1,1,2)$ : three-parameter case of GE distribution with location parameter is equal to 1, scale parameter is equal to 1 and shape parameter is equal to 2.
- 3. Sample from  $G(0,2,2)$ : gamma distribution with location parameter is equal to 0, scale parameter is equal to 2 and shape parameter is equal to 2.

4. Sample from GE(2,2,2): gamma distribution with location parameter is equal to 2, scale parameter is equal to 2 and shape parameter is equal to 2.

The above four order statistics samples with the analogous quantiles of order statistics from  $GE(0,1,2)$  are used to run the test. The results of the test at 5% significance level and at  $\alpha=2$  (whether accept (A) or reject (R)  $H_0$ ) are given for different values in the following table.

**Table 5.** Results at 5% significance,  $\alpha = 2$ 

<b>Decision</b>							
GE(0,1,2)	GE(1,1,2)	G(0,2,2)	G(2,2,2)				

## **Conclusions**

This article proposed a test formula parallel to the one developed by Sultan [\(2007\)](#page-8-0). It was found to be simple and can be used for any sample size. Moreover, it is more effective with respect to power evaluation and coverage probabilities.

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