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A Comparison of Shape and Scale Estimators of the Two-Parameter Weibull Distribution

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Weibull distributions are widely used in reliability and survival analysis. In this paper, different methods to estimate the shape and scale parameters of the two-parameter Weibull distribution have been reviewed and compared, based on the bias, mean square error and variance. Because a theoretical comparison is not possible, an extensive simulation study has been conducted to compare the performance of different estimators. Based on the simulation study it was observed that MLE consistently performs better than other methods.

Keywords: Two-parameter Weibull distribution, scale parameters, shape parameters

Introduction

The Weibull distribution is a commonly used model in reliability, life time and environmental data analysis. A considerable literature discussing the methods of estimation of Weibull parameters exists (Sharoon, et al., 2012; Saralees et al., 2011; Saralees et al., 2008) because of its applications in different fields. Kantar and Senoglu (2008) did a simulation comparison of different estimators for scale parameter when shape is known. Balakrishnan and Kateri (2008) showed the existence and uniqueness of maximum likelihood estimates (MLE) of Weibull distribution. Dubey (1967) derived the percentile estimators (Percentile 1) which uses 4 different percentiles to estimate the shape and scale parameters. Seki and Yokoyama (1993) proposed a simple and robust method that uses only two percentiles, 31st and 63rd percentile (Percentile 2) to estimate both parameters. Moment estimators (MOM) and median rank regression estimators (MRRS) are also commonly used in literature (Kantar and Senoglu, 2008) because of their

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easiness in computation. Existing methods (namely MLE, MOM, MRRS, Percentile 1, and Percentile 2) for estimating both shape and scale parameters of two-parameter Weibull distribution are here reviewed and compared. A simulation study has been conducted to compare the performance of these methods under same simulation conditions.

Statistical Methodology

The Weibull distribution has the probability density function,

$$f(x) = \alpha\beta^{-\alpha} x^{\alpha-1} e^{-\left(\frac{x}{\beta}\right)^\alpha}$$
 for $x \geq 0, \alpha > 0, \beta > 0$, where α is the shape parameter and β is the scale parameter. The cumulative distribution function is given by

$$F(x) = 1 - e^{-\left(\frac{x}{\beta}\right)^\alpha} \text{ for } x \geq 0.$$

The distribution is reversed J-shaped when $\alpha < 1$, exponential when $\alpha = 1$ and bell-shaped when $\alpha > 1$ (Kantar and Senoglu, 2008). Because of its wide-variety of shapes it is used extensively in practice for modeling real life data in different fields.

Maximum Likelihood Estimators (MLE)

The log-likelihood function of a random sample from the two-parameter Weibull distribution is given by

$$\ln L = n \ln \alpha - n \alpha \ln \beta + (\alpha - 1) \sum \ln x - \sum \left(\frac{x}{\beta} \right)^\alpha.$$

This will yield the following two score equations

$$\begin{aligned} \frac{\partial \ln L}{\partial \beta} &= -\frac{n\alpha}{\beta} + \frac{\alpha}{\beta^{\alpha+1}} \sum x^\alpha = 0 \text{ and} \\ \frac{\partial \ln L}{\partial \alpha} &= -\frac{n}{\alpha} - n \ln \beta + \sum \ln x + \frac{\ln \beta}{\beta^\alpha} \sum x^\alpha - \frac{1}{\beta^\alpha} \sum x^\alpha \ln x = 0 \end{aligned}$$

The above two equations can be solved numerically to obtain MLEs.

Moment Estimators (MOM)

The moment estimators are obtained by equating the population moments to the corresponding sample moments. The first and second moments of Weibull distribution are respectively

$$\mu'_1 = \beta \Gamma(1 + \alpha^{-1}) \text{ and}$$

$$\mu'_2 = \beta^2 \Gamma(1 + 2\alpha^{-1})$$

The first two moments from the sample are $m'_1 = \frac{1}{n} \sum x$ and $m'_2 = \frac{1}{n} \sum x^2$.

The moment estimates are obtained by solving the following two equations

$$m'_1 = \beta \Gamma(1 + \alpha^{-1})$$

$$m'_2 = \beta^2 \Gamma(1 + 2\alpha^{-1})$$

Median Rank Regression Estimators (MRRS)

MRR is a procedure for estimating the Weibull parameters by fitting a least squares regression line through the points on a probability plot. Thus,

$$\log(1 - F(x)) = -\left(\frac{x}{\beta}\right)^\alpha \text{ and hence}$$

$$\log(\log(1 - F(x))) = \alpha \log x = \alpha \log \beta.$$

This is now a linear model and method of least squares can be used to estimate α and β . The sample data are first sorted in ascending order and then following Abernethy (2006), the distribution function, $F(x_i)$ is approximated for each point (x_i) in the sorted sample as $F(x_i) = \frac{i - 0.3}{n + 0.4}$, where I is the ascending rank of the data point x_i .

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Percentile Estimators (Percentile 1)

Percentile estimators for both shape and scale parameters were derived by Dubey (1967). He proposed an estimator based on 17th and 97th percentiles for shape parameter and one based on 40th and 82nd percentile for scale parameter. The formulae for the shape and scale percentile estimators are presented here; for details refer to Dubey (1967). Let $p_1 = 0.1673$ and $p_2 = 0.9737$. Define $k_1 = \log(-\log(1 - p_1)) - \log(-\log(1 - p_2))$. Let y_1 and y_2 represent the 100 p_1 th percentile from the data. Then

$$\hat{\alpha} = \frac{-k_1}{\log(y_1) - \log(y_2)}$$

Similarly to estimate β , define $p_3 = 0.3978$ and $p_4 = 0.8211$. Let $k_2 = \log(-\log(1 - p_3)) - \log(-\log(1 - p_4))$; $k_3 = -\log(1 - p_3)$ and $w = 1 - \frac{\log(k_3)}{k_2}$. Let y_3 and y_4 represent the 100 p_3 th and 100 p_4 th percentile from the data. Then

$$\beta = \exp(w \log(y_3) + (1-w) \log(y_4)).$$

Improved Percentile Estimators (Percentile 2)

Seki and Yokoyama (1993) proposed this simple and robust method that uses only two percentiles, 31st and 63rd percentile to estimate α and β . The Weibull cumulative distribution function is given by

$$F(x) = 1 - e^{-\left(\frac{x}{\beta}\right)^\alpha} \quad \text{for } x \geq 0.$$

Hence the 100 p th percentile of the Weibull distribution can be written as $x_p = \beta(-\log(1 - p))^{1/\beta}$. Then the 100(1 - e^{-1}) = 63.2th percentile is $x_{0.632} = \beta$ for any Weibull distribution. This can be used to compute $\hat{\beta}$. Therefore, the estimate of the shape parameter can be obtained as $\hat{\alpha} = \frac{\log(-\log(1 - p))}{\log\left(\frac{x_p}{x_{0.632}}\right)}$. Seki and

Yokoyama (1993) approximated the numerator of this estimator as -1 and then obtained $p = 0.31$, approximately, to obtain $\hat{\alpha}$.

Simulation Study

A simulation study has been conducted to explore the performances of the different methods discussed in this article.

Simulation Technique

The main objective of this study is to compare the performance of five different methods to estimate the shape and scale parameters of two-parameter Weibull distribution. Weibull distribution with parameters scale = 10 and shape = 0.5, 1, 1.5, 2, 3 and 4 were used to generate 5,000 samples of sizes n = 5, 10, 20, 30, 50 and 100. The estimates are compared using the values of average bias, mean squared error (MSE) and variance. The simulation was done using statistical software R version 2.15.2.

Results and Discussion

The results of the simulation are shown in [Tables 1 to 3](#). The bias and MSEs from Weibull (10, 0.5) and Weibull (10, 3) are also presented in [Figures 1 to 4](#). From [Tables 1 to 3](#), it can be observed that as sample size increases, bias, MSE and variance decrease. For small sample size, the performance of methods differs significantly. For all methods, absolute bias, MSE and variance decrease as sample size increases. It can be observed from [Tables 1 to 3](#) and [Figures 1 to 4](#), in almost all cases MLE performed better than the other 4 methods and percentile method-1 performed the worst. In some situations, MRRS also performs well, especially for shape estimates. It can also be observed that both percentile estimators perform poorly in estimation of shape. There is no consistency in the performance of estimates by the method of moments. Because MLE is performing consistently better than the other 4 methods practitioners are encouraged to use MLE whenever possible.

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Table 1. Bias, Variance and MSE of both Scale and Shape estimates $\alpha=10$ and $\beta=0.5; 1$

α, β	n	Scale					Shape					
		MLE	MOM	MRRS	Perctle1	Perctle2	MLE	MOM	MRRS	Perctle1	Perctle2	
10, 0.5	5	Bias	2.929	6.216	5.361	5.340	4.394	0.219	0.351	0.025	0.661	0.497
		Vars	156.930	224.302	219.866	282.068	197.209	0.147	0.092	0.085	5.231	0.313
		MSEs	165.511	262.939	248.610	310.581	216.516	0.195	0.216	0.085	5.668	0.560
	10	Bias	1.525	4.606	2.998	2.334	1.773	0.085	0.218	-0.015	0.283	0.193
		Vars	65.147	96.815	84.821	110.433	73.834	0.031	0.033	0.026	0.349	0.050
		MSEs	67.473	118.026	93.811	115.879	76.977	0.038	0.081	0.026	0.429	0.087
	20	Bias	0.705	3.217	1.643	0.979	0.656	0.036	0.136	-0.021	0.112	0.079
		Vars	25.634	40.337	32.389	43.623	30.013	0.010	0.016	0.012	0.069	0.015
		MSEs	26.131	50.689	35.088	44.582	30.444	0.012	0.035	0.013	0.082	0.021
10, 1	30	Bias	0.556	2.621	1.258	0.707	0.442	0.026	0.107	-0.017	0.072	0.054
		Vars	17.111	25.216	20.965	26.613	19.722	0.006	0.013	0.008	0.036	0.009
		MSEs	17.420	32.084	22.548	27.113	19.917	0.007	0.024	0.008	0.041	0.012
	50	Bias	0.283	1.912	0.783	0.405	0.207	0.014	0.076	-0.016	0.039	0.034
		Vars	9.585	15.077	11.172	15.220	11.347	0.003	0.008	0.005	0.017	0.005
		MSEs	9.665	18.731	11.785	15.384	11.390	0.004	0.014	0.005	0.019	0.006
	100	Bias	0.174	1.257	0.487	0.236	0.158	0.008	0.048	-0.011	0.021	0.018
		Vars	4.664	8.153	5.320	7.373	5.727	0.002	0.005	0.003	0.007	0.002
		MSEs	4.694	9.734	5.557	7.429	5.752	0.002	0.008	0.003	0.008	0.003
10, 1	5	Bias	0.284	0.312	1.232	0.485	0.304	0.443	0.353	0.049	1.297	0.926
		Vars	23.380	23.097	27.835	29.087	24.408	0.621	0.411	0.347	10.932	1.147
		MSEs	23.461	23.194	29.353	29.322	24.501	0.817	0.535	0.349	12.614	2.005
	10	Bias	0.190	0.222	0.840	0.223	0.060	0.173	0.173	-0.032	0.507	0.371
		Vars	11.082	11.192	12.944	15.822	12.013	0.127	0.108	0.106	0.991	0.199
		MSEs	11.118	11.241	13.650	15.872	12.016	0.157	0.137	0.107	1.247	0.336
	20	Bias	0.072	0.100	0.486	0.067	-0.047	0.074	0.087	-0.042	0.225	0.158
		Vars	5.576	5.707	6.335	8.386	6.348	0.043	0.046	0.048	0.281	0.063
		MSEs	5.581	5.717	6.571	8.390	6.350	0.049	0.053	0.050	0.331	0.088
10, 1	30	Bias	0.075	0.098	0.412	0.049	-0.006	0.049	0.062	-0.040	0.141	0.108
		Vars	3.671	3.796	4.182	5.412	4.340	0.025	0.029	0.032	0.134	0.037
		MSEs	3.676	3.805	4.351	5.414	4.340	0.027	0.033	0.033	0.154	0.048
	50	Bias	0.050	0.062	0.281	0.028	-0.010	0.029	0.039	-0.031	0.087	0.069
		Vars	2.173	2.269	2.426	3.332	2.635	0.014	0.018	0.020	0.072	0.020
		MSEs	2.175	2.273	2.504	3.332	2.635	0.015	0.019	0.021	0.080	0.025
	100	Bias	0.032	0.034	0.179	0.012	0.002	0.016	0.021	-0.021	0.045	0.037
		Vars	1.117	1.178	1.232	1.698	1.349	0.007	0.010	0.010	0.031	0.010
		MSEs	1.118	1.179	1.264	1.698	1.349	0.007	0.010	0.011	0.033	0.011

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Table 2. Bias, Variance and MSE of both Scale and Shape estimates $\alpha=10$ and $\beta=1.5$; 2

α, β	n	Scale					Shape				
		MLE	MOM	MRRS	Perctle1	Perctle2	MLE	MOM	MRRS	Perctle1	Perctle2
10, 1.5	Bias	-0.081	-0.149	0.513	-0.149	-0.226	0.650	0.398	0.059	2.045	1.298
	5 Vars	9.333	9.331	10.477	11.069	9.546	1.374	0.947	0.713	47.203	2.365
	MSEs	9.340	9.353	10.740	11.091	9.597	1.797	1.106	0.717	51.386	4.050
	Bias	0.038	-0.002	0.451	-0.016	-0.092	0.257	0.171	-0.048	0.755	0.537
	10 Vars	4.948	4.986	5.542	6.870	5.493	0.275	0.231	0.232	2.331	0.427
	MSEs	4.950	4.986	5.746	6.870	5.501	0.342	0.260	0.235	2.901	0.715
	Bias	0.039	0.016	0.317	0.016	-0.046	0.114	0.081	-0.065	0.339	0.248
	20 Vars	2.508	2.517	2.763	3.802	2.851	0.097	0.090	0.109	0.606	0.144
	MSEs	2.509	2.517	2.863	3.802	2.853	0.110	0.096	0.113	0.721	0.205
10, 2	Bias	0.014	-0.003	0.227	-0.006	-0.064	0.075	0.054	-0.054	0.209	0.167
	30 Vars	1.669	1.677	1.816	2.439	1.947	0.055	0.054	0.074	0.324	0.083
	MSEs	1.669	1.677	1.867	2.439	1.952	0.061	0.057	0.077	0.368	0.111
	Bias	-0.008	-0.019	0.145	-0.021	-0.050	0.041	0.030	-0.050	0.121	0.099
	50 Vars	0.973	0.979	1.048	1.486	1.166	0.031	0.032	0.045	0.155	0.048
	MSEs	0.973	0.979	1.069	1.486	1.168	0.033	0.033	0.048	0.170	0.058
	Bias	0.003	-0.004	0.102	-0.020	-0.027	0.022	0.016	-0.032	0.067	0.054
	100 Vars	0.493	0.495	0.533	0.748	0.591	0.014	0.016	0.023	0.070	0.021
	MSEs	0.493	0.495	0.543	0.749	0.592	0.015	0.016	0.024	0.074	0.024
5, 2	Bias	-0.102	-0.123	0.339	-0.200	-0.263	0.868	0.510	0.089	2.499	1.673
	5 Vars	5.328	5.376	5.763	6.306	5.441	2.301	1.833	1.321	43.582	3.716
	MSEs	5.339	5.391	5.878	6.345	5.511	3.055	2.093	1.329	49.828	6.516
	Bias	-0.049	-0.065	0.254	-0.132	-0.189	0.338	0.191	-0.060	1.024	0.699
	10 Vars	2.797	2.830	3.090	3.811	3.062	0.459	0.390	0.394	4.718	0.699
	MSEs	2.799	2.834	3.154	3.829	3.098	0.573	0.427	0.398	5.767	1.187
	Bias	-0.041	-0.051	0.157	-0.089	-0.136	0.151	0.087	-0.080	0.473	0.331
	20 Vars	1.359	1.368	1.461	2.068	1.560	0.171	0.159	0.197	1.296	0.262
	MSEs	1.361	1.371	1.486	2.076	1.578	0.194	0.166	0.203	1.520	0.371
10, 1	Bias	-0.004	-0.009	0.157	-0.023	-0.059	0.096	0.054	-0.081	0.275	0.222
	30 Vars	0.911	0.915	0.991	1.347	1.059	0.096	0.090	0.125	0.540	0.142
	MSEs	0.911	0.915	1.016	1.348	1.062	0.105	0.093	0.131	0.615	0.191
	Bias	-0.007	-0.010	0.110	-0.028	-0.049	0.059	0.035	-0.063	0.166	0.141
	50 Vars	0.543	0.546	0.596	0.829	0.654	0.055	0.054	0.080	0.269	0.084
	MSEs	0.543	0.546	0.608	0.830	0.656	0.059	0.055	0.084	0.297	0.104
	Bias	0.001	-0.002	0.071	-0.020	-0.022	0.031	0.019	-0.040	0.088	0.073
	100 Vars	0.276	0.276	0.296	0.427	0.332	0.025	0.025	0.041	0.120	0.037
	MSEs	0.276	0.276	0.301	0.427	0.333	0.026	0.026	0.042	0.128	0.043

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Table 3. Bias, Variance and MSE of both Scale and Shape estimates $\alpha=10$ and $\beta=3; 4$

α, β	n	Scale					Shape				
		MLE	MOM	MRRS	Perctle1	Perctle2	MLE	MOM	MRRS	Perctle1	Perctle2
10, 3	Bias	-0.166	-0.138	0.122	-0.275	-0.320	1.295	0.767	0.131	3.899	2.417
	5 Vars	2.392	2.439	2.515	2.865	2.489	5.768	5.041	3.285	243.739	8.976
	MSEs	2.419	2.458	2.530	2.941	2.591	7.446	5.629	3.302	258.938	14.818
	Bias	-0.053	-0.037	0.147	-0.129	-0.152	0.505	0.276	-0.102	1.461	1.048
	10 Vars	1.221	1.234	1.290	1.659	1.352	1.100	0.998	0.911	9.896	1.722
	MSEs	1.224	1.235	1.312	1.676	1.374	1.354	1.074	0.922	12.032	2.820
	Bias	-0.017	-0.009	0.120	-0.060	-0.070	0.218	0.113	-0.134	0.689	0.487
	20 Vars	0.602	0.605	0.649	0.896	0.690	0.374	0.359	0.428	2.594	0.572
	MSEs	0.602	0.605	0.663	0.900	0.695	0.422	0.372	0.445	3.070	0.810
10, 4	Bias	-0.022	-0.017	0.082	-0.049	-0.057	0.153	0.084	-0.110	0.452	0.341
	30 Vars	0.415	0.417	0.442	0.618	0.487	0.230	0.227	0.294	1.387	0.343
	MSEs	0.416	0.417	0.448	0.620	0.491	0.253	0.234	0.306	1.592	0.459
	Bias	-0.020	-0.017	0.058	-0.039	-0.044	0.083	0.044	-0.101	0.254	0.209
	50 Vars	0.243	0.243	0.259	0.358	0.288	0.121	0.123	0.185	0.643	0.183
	MSEs	0.243	0.243	0.262	0.360	0.290	0.128	0.125	0.196	0.707	0.226
	Bias	0.000	0.001	0.047	-0.003	-0.014	0.035	0.014	-0.075	0.098	0.096
	100 Vars	0.118	0.118	0.128	0.186	0.144	0.058	0.059	0.092	0.265	0.087
	MSEs	0.118	0.118	0.130	0.186	0.144	0.059	0.059	0.098	0.274	0.096
10, 4	Bias	-0.125	-0.081	0.091	-0.230	-0.258	1.716	1.051	0.164	4.996	3.156
	5 Vars	1.388	1.411	1.442	1.683	1.471	9.001	7.677	5.003	141.435	13.999
	MSEs	1.404	1.418	1.451	1.736	1.537	11.945	8.782	5.029	166.393	23.962
	Bias	-0.078	-0.054	0.076	-0.140	-0.156	0.686	0.401	-0.131	2.017	1.402
	10 Vars	0.670	0.675	0.709	0.913	0.747	2.063	1.999	1.723	16.360	3.181
	MSEs	0.676	0.678	0.714	0.932	0.772	2.535	2.159	1.740	20.429	5.146
	Bias	-0.033	-0.021	0.070	-0.056	-0.076	0.305	0.166	-0.174	0.875	0.659
	20 Vars	0.362	0.364	0.391	0.545	0.420	0.691	0.695	0.774	4.648	1.023
	MSEs	0.363	0.365	0.396	0.548	0.426	0.783	0.723	0.804	5.414	1.457
10, 4	Bias	-0.024	-0.016	0.055	-0.038	-0.055	0.194	0.107	-0.156	0.549	0.444
	30 Vars	0.229	0.230	0.245	0.339	0.271	0.388	0.411	0.523	2.300	0.567
	MSEs	0.229	0.230	0.248	0.341	0.274	0.425	0.423	0.547	2.601	0.764
	Bias	-0.016	-0.011	0.042	-0.025	-0.035	0.112	0.060	-0.132	0.322	0.271
	50 Vars	0.138	0.139	0.149	0.209	0.163	0.224	0.237	0.323	1.087	0.342
	MSEs	0.138	0.139	0.151	0.209	0.165	0.237	0.241	0.341	1.190	0.415
	Bias	-0.013	-0.011	0.023	-0.021	-0.025	0.060	0.036	-0.087	0.156	0.147
	100 Vars	0.066	0.067	0.071	0.104	0.082	0.108	0.116	0.174	0.478	0.157
	MSEs	0.066	0.067	0.072	0.105	0.083	0.111	0.118	0.181	0.503	0.178

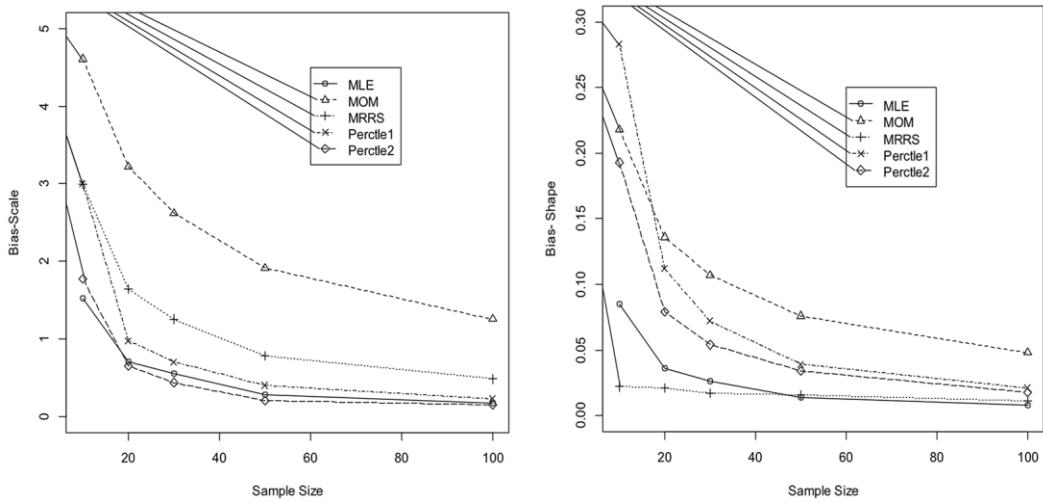


Figure 1. Absolute Bias of Scale parameter estimate (left), Shape parameter estimate (right) vs. Sample size from Weibull (10, 0.5)

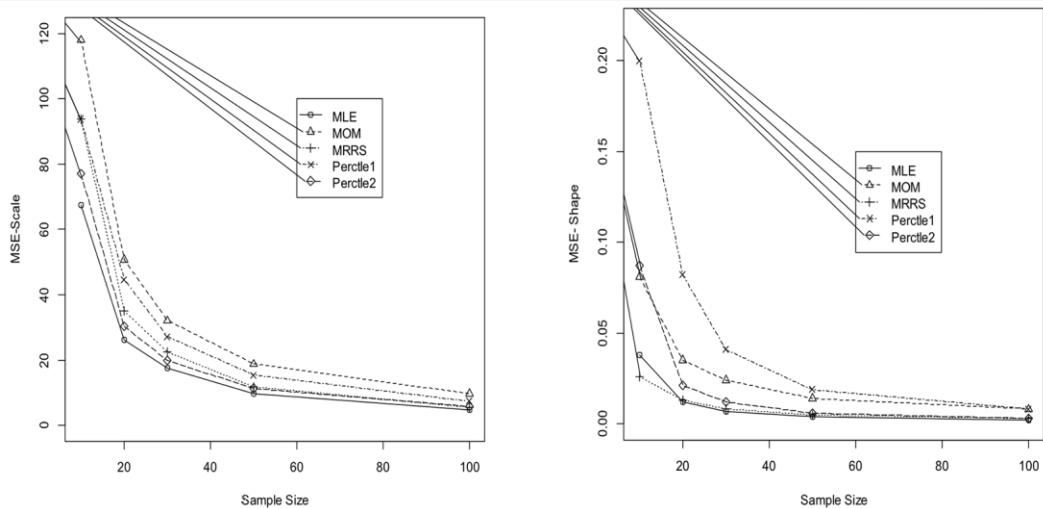


Figure 2. MSE of Scale parameter estimate (left), Shape parameter estimate (right) vs. Sample size from Weibull (10, 0.5)

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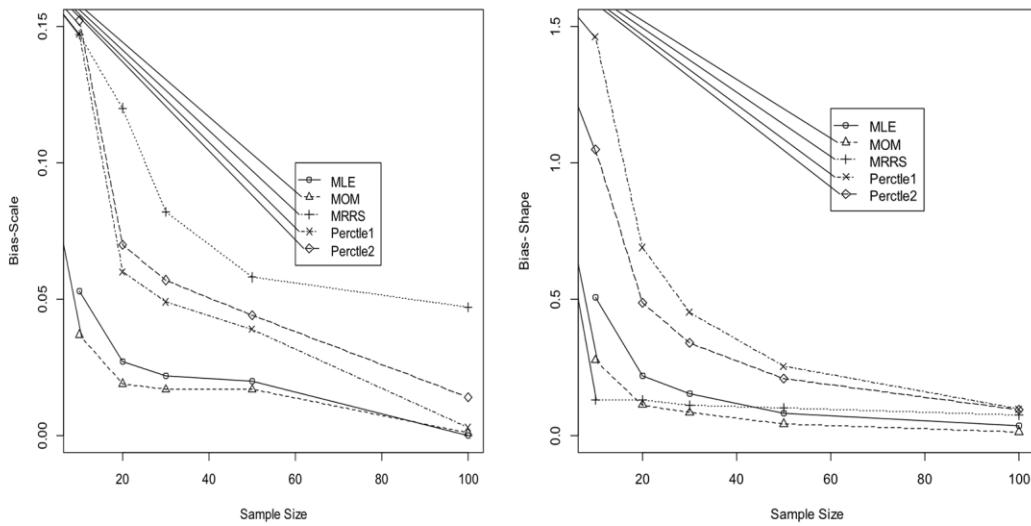


Figure 3. Absolute Bias of Scale parameter estimate (left), Shape parameter estimate (right) vs. Sample size from Weibull (10, 3)

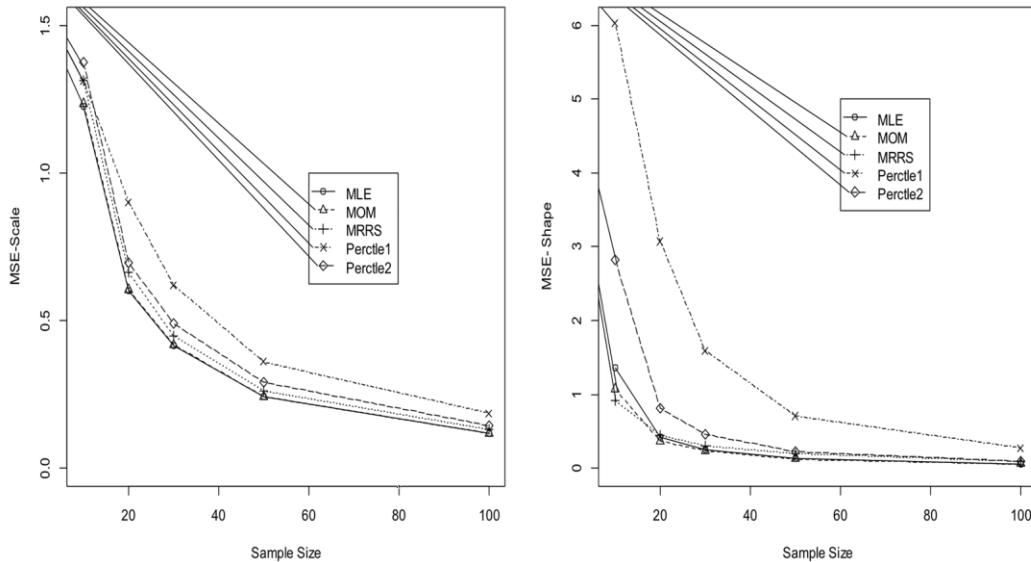


Figure 4. MSE of Scale parameter estimate (left), Shape parameter estimate (right) vs. Sample size from Weibull (10, 3)

Example

Many researchers modeled wind data using the Weibull distribution (Dorvlo, 2002; Weisser, 2003; Celik, 2003). The five methods with an example in Battacharya and Bhattacharjee (2010) will be discussed next. This example provides the average monthly wind speed (m/s) of Kolkata from 1st March 2009 to 31st March 2009. [Table 4](#) presents the data set.

The estimates of the two-parameter Weibull distribution obtained by fitting to the data using the methods discussed in the article are given in the [Table 5](#). It look like the Percentile 1 and Percentile 2 estimates are at the extreme ends and the MLE estimates lie somewhat between the values of other estimates.

Table 4. Average daily wind speed in Kolkata during March 2009.

Date	Speed (m/s)	Date	Speed (m/s)
1	0.56	17	0.28
2	0.28	18	0.83
3	0.56	19	1.39
4	0.56	20	1.11
5	1.11	21	1.11
6	0.83	22	0.83
7	1.11	23	0.56
8	1.94	24	0.83
9	1.11	25	1.67
10	0.83	26	1.94
11	1.11	27	1.39
12	1.39	28	0.83
13	0.28	29	2.22
14	0.56	30	1.67
15	0.28	31	2.22
16	0.28		

Table 5. Estimates of two-parameter Weibull by different methods.

Method	Scale	Shape
MLE	1.1550	1.9081
MOM	1.1501	1.8456
MRRS	1.1636	1.8031
Percentile 1	1.1100	1.8055
Percentile 2	1.1816	2.1704

Conclusion

Five different methods for the joint estimation of both scale and shape parameters of two-parameter Weibull distribution were reviewed in this article. A simulation study was conducted to compare the five methods based on bias, mean square error and variance of estimates. From simulation results, it was observed that MLE performs consistently better than MOM, MRRS, percentile method and

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improved percentile method and therefore MLE estimates are recommended to the practitioners.

References

- Abernethy, R. B. (2006). (5th Ed.). *The new Weibull handbook : reliability & statistical analysis for predicting life, safety, risk, support costs, failures, and forecasting warranty claims, substantiation and accelerated testing, using Weibull, Log normal, crow-AMSAA, probit, and Kaplan-Meier models*. North Palm Beach, FL: R. B. Abernethy.
- Abernethy, R. B., Breneman, J. E., Medlin, C. H., & Reinman, G. L. (1983). *Weibull Analysis Handbook*. Air Force Wright Aeronautical Laboratories Technical Report AFWAL-TR- 83-2079. Available at <http://handle.dtic.mil/100.2/ADA143100>.
- Balakrishnan, N. & Kateri, M. (2008). On the maximum likelihood estimation of parameters of Weibull distribution based on complete and censored data. *Statistics & Probability Letters*, 78(17): 2971–2975.
- Battacharya, P., & Bhattacharjee, R. (2010). A study on Weibull distribution for estimating the parameters. *Journal of Applied Quantitative Methods*, 5(2): 234-241.
- Celik, A. N. (2003). Energy output estimation for small-scale wind power generators using Weibull-representative wind data. *Journal of Wind Engineering and Industrial Aerodynamics*, 91: 693-707. doi: 10.1016/S0167-6105(02)00471-3
- Dorvlo, A. S. (2002). Estimating wind speed distribution. *Energy Conservation and Management*, 43: 2311-2318.
- Dubey, S. Y. D. (1967). Normal and Weibull distributions. *Naval Research Logistics Quarterly*, 14(1): 69–79. doi: 10.1002/nav.3800140107
- Kantar, Y. M. & Senoglu, B. (2008). A Comparative Study for the Location and Scale Parameters of the Weibull Distribution with a given Shape Parameter. *Computers & Geosciences*, 34: 1900-1909.
- Saralees, N., & Firoozeh, H. (2011). An extension of the exponential distribution. *Statistics - A Journal of Theoretical and Applied Statistics*, 45(6): 543-558.
- Saralees, N., & Kotz, S. (2008). Strength modeling using Weibull distributions. *Journal of Mechanical Science and Technology*, 22: 1247-1254.

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- Seki, T., & Yokoyama, S. (1993). Simple and robust estimation of the Weibull parameters. *Microelectronics Reliability*, 33(1): 45-52.
- Sharoon, H., Muhammad, Q. S., Muhammad, M., & Kibria, B. M. G. (2012). A Note on Beta Inverse - Weibull Distribution. *Communication in Statistics - Theory and Methods*, 42(2): 320-335.
- Weisser, D. (2003). A wind energy analysis of Grenada: an estimation using the ‘Weibull’ density function. *Renewable Energy*; 28(11): 1803-1812.